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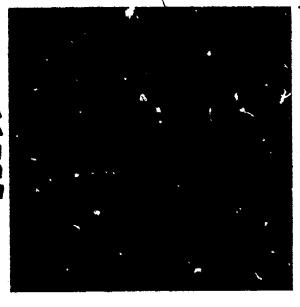
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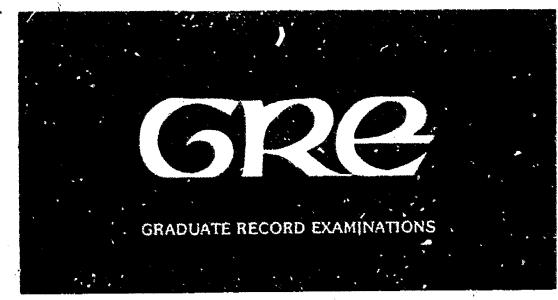
ABSTRACT

This Graduate Record Examination (GRE) study assesses: (1) the relative contribution of a vocabulary score (consisting of GRE General Test antonyms and analogies) and a reading comprehension score (consisting of GRE sentence completion and reading comprehension sets) to the prediction of self-reported undergraduate grade point average (GPA); and (2) criterion-related validity patterns for item-type part scores on the GRE quantitative and analytical measures. Data from GRE files for 9,375 examinees in 12 fields of study representing 437 undergraduate departments from 149 colleges and universities were standardized within each undergraduate department, and then pooled for analysis by field. There were differences by major field in average persormance on the various item-type part scores within each test. The reading comprehension subtest carried most of the predictive load in the GRE verbal measure. Item-type part scores on the other measures also exhibited differential patterns of relationships with the self-reported undergraduate grade point average. The findings suggest that the different item types within the respective broad ability measures may be tapping somewhat unique skills and abilities and that further exploration of their potential contribution is in order. (Author/BS)

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THE RELATIONSHIP OF GRE GENERAL TEST ITEM-TYPE

PART SCORES TO UNDERGRADUATE GRADES

Kenneth M. Wilson

GRE Board Professional Report GREB No. 81-22P ETS Research Report 84-38

February, 1985

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EDUCATIONAL TESTING SERVICE, PRINCETON, NJ

Abstract

This study was undertaken (a) to assess the relative contribution of a vocabulary score made up of GRE General Test antonyms and analogies and a reading comprehension score made up of GRE sentence completions and reading comprehension sets to prediction of an academic criterion (self-reported undergraduate grade point average) and (b) to assess patterns of criterion-related validity for item-type part scores on the GRE quantitative and analytical measures as well.

The study was based on data from GRE files for 9,375 examinees in 12 fields of study representing 437 undergraduate departments from 149 colleges and universities. All data were standardized within each undergraduate department and then pooled for analysis by field.

There were differences by major field in average performance on the various item-type part scores within each test. The reading comprehension subtest was found to carry most of the predictive load in the GRE verbal measure (consistent with findings for the reading comprehension subscore on the SAT verbal measure). Item-type part scores on the other measures also exhibited differential patterns of relationships with the self-reported undergraduate grade point average.

The findings suggest that the different item types within the respective broad ability measures may be tapping somewhat unique skills and abilities and that further exploration of their potential contribution is in order.



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The Relationship of GRE General Test Item-Type Part Scores to Undergraduate Grades

Kenneth M. Wilson Educational Testing Service

Introduction

The GRE General (Aptitude) Test provides measures of developed verbal, quantitative, and analytical abilities.* Only total verbal, quantitative, and analytical scores are reported. However, the three measures include different types of items that are thought of as being different methods of measuring their respective constructs (Rock, Werts, & Grandy, 1982).

The verbal measure employs four types of questions or items: antonyms, analogies, sentence completions, and reading comprehension sets designed to test the ability to identify (a) words that are opposite in meaning, (b) words or phrases that are related to each other in the same way as other words or phrases, and (c) words that are logically and stylistically consistent with the sentence in which they appear; and (d) the ability to recognize in a reading passage the main ideas, information explicitly provided, implied ideas, the attitude of the author, and the like.

Three item types employed in the quantitative measure: are quantitative comparisons (testing the ability to reason quickly and accurately regarding the relative sizes of two quantities or to perceive. that not enough information is available to make such a decision); discrete quantitative items measuring basic mathematical skills or mathematics (balanced among question requiring arithmetic, algebra, and geometry and designed to test basic mathematical skills and understandings of concepts, at levels applicable to individuals who have not specialized in mathematics); and data interpretation (testing the ability to synthesize information presented in tabular or graphic form, to select data appropriate for answering a question, and so on).

The 1981 revision of the analytical measure includes two item types: analytical reasoning items (testing the ability to understand a given structure of arbitrary relationships among fictitious entities, deduce new information from given relationships, and the like); and logical reasoning items (testing the ability to understand, analyze, and evaluate arguments, recognize the point of an argument or the assumptions on which it is based, analyze evidence, and the like).

Although a continuing effort is made to obtain empirical evidence regarding the validity of the total verbal, quantitative, and analytical

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^{*}For detailed descriptions of tests and item types, see, for example, ETS (1981). In October 1977, a restructured version of the GRE General Test including a newly developed analytical ability measure was introduced. Evidence of its predictive validity with respect to graduate grades was obtained in a cooperative study (Wilson, 1982). However, internal research indicated the need for some change in the item content of the 1977 analytical measure and, in October 1981, a revised analytical measure was introduced. See Wild, Swinton, and Wallmark (1982) for a review of factors involved in the 1981 revision.

scores for predicting perfo mance in graduate study, little attention has been given to study of the predictive validity and diagnostic potential of part scores based on the various item types—in large part because of the lack of any compelling a priori evidential or theoretical basis for expecting differential predictive validity for part scores based on different item types measuring more general basic constructs such as verbal or quantitative ability.

For example, items regardless of type are selected on the basis of internal consistency criteria designed among other things to assure the comparative homogeneity of the respective ability measures. This is conducive to relatively high intercorrelations among items and between individual items and the total scores on the respective tests. Such conditions theoretically militate against the likelihood, for example, that predictions based on regression-weighted composites of part scores would be consistently better than predictions based on the total score (in which the potential item-type part scores are weighted roughly according to their length). Although factor analytic studies (for example, Powers & Swinton, 1981; Rock, Werts, & Grandy, 1982) have suggested that word knowledge (vocabulary) and reading items (reading comprehension) - e distinguishable factorially, this evidence alone has not been sufficiently persuasive to suggest that predictions based on the "vocabulary" items and predictions based on "reading comprehension" items would be very different.

However, the need for an empirical evaluation of the predictive validity of item-type part scores on the GRE General Test was indicated by the results of undergraduate-level validity studies involving verbal item-type part scores on the College Board Scholastic Aptitude Test (SAT). For several years, vocabulary (VO) and reading comprehension (RC) scores have been reported in addition to the total SAT verbal score. The vocabulary score is based on antonyms and analogies and the reading comprehension score on sentence completions and reading comprehension sets. These items are completely parallel in type to those included in the GRE verbal measure.

Based on internal analyses of the results of 110 studies conducted by the College Board Validity Study Service (VSS) at ETS (Ramist, 1981a; 1981b) in which colleges had specified vocabulary, reading comprehension, and total SAT verbal scores as predictors of freshman grades, the following findings emerged:

- o The average validity of the reading comprehension score alone (.373) was only .003 points lower than that for the entire verbal score (.376).
- o In almost one-half of the samples studied, the observed validity of the reading comprehension score was actually greater than that for the SAT verbal score, including the vocabulary score, the validity of which was consistently lower than that of the reading comprehension score.
- o When vocabulary and reading comprehension scores were combined in regression-weighted composites, the vocabulary score in a number of instances was negatively weighted, although its simple correlation with the GPA criterion was positive, indicating suppression of vocabulary



variance in reading comprehension—that is, suggesting that the criterion—related variance in the vocabulary measure was being tapped sufficiently by the reading comprehension measure with which the vocabulary score is substantially correlated.

o There was little improvement in predicting freshman grade point average when separate vocabulary and reading comprehension scores replaced the SAT total verbal score in regression equations including SAT mathematical scores and the high school record.

These results were inconsistent with expectation and raised questions regarding the relative predictive role of the SAT vocabulary and reading comprehension items.* The present study was undertaken to assess the relationship to academic performance of similarly constructed GRE vocabulary and reading comprehension item-type part scores (and of item-type part scores based on items in the quantitative and analytical tests as well).

Study Design, Sample, and Procedures

The academic performance criterion selected for this exploratory study was self-reported undergraduate grade point average (SR-UGPA) routinely supplied by most GRE examinees during the process of test-registration.**

The SR-UGPA has been found to be a useful research surrogate for an officially computed UGPA as a predictor of graduate GPA (Wilson, 1982). Moreover, patterns of coefficients for GRE verbal, quantitative, and analytical scores vs SR-UGPA, computed for samples of undergraduate students majoring in selected fields (for example, Miller & Wild, 1979) appear to be similar to patterns of coefficients for these predictors vs graduate GPA (for example, Wilson, 1982).

It was reasoned that results of an exploratory study involving SR-UGPA as the academic performance criterion would provide a useful empirical basis for initial assessment of the validity of item-type part scores. Such a study would also contribute to further understanding of the utility of the SR-UGPA in research concerned with test validation.



^{*}Several lines of inquiry have been initiated, including a study of the relationship of vocabulary and reading comprehension scores to self-reported high school rank, a study of the statistical properties of the four item-types included in the SAT verbal measure, and a study of the criterion-related validity of specific verbal item types on one form of the SAT verbal test (Schrader, 1984).

^{**}Examinees are asked to report UGPA in the major field and UGPA over the last two college years. The criterion employed was the average of the two self-reported undergraduate grade point averages.

The study was designed to simulate conditions characteristic of graduate-level validity studies in which comparable data sets for several small departmental samples are pooled for analysis by field or discipline (for example, Wilson, 1979; 1982).

Study Sample and Data

The study sample and basic study data were taken from GRE files on examinees tested between Oct. 1, 1981, and Sept. 30, 1982. The study sample included only examinees who reported better communication in English than in any other language, who were tested as enrolled undergraduates of nonenrolled college graduates no more than two years beyond the bachelor's degree, and who named both a field of study and an undergraduate school. Following procedures described below, data were obtained for examinees representing both (a) a relatively large number of undergraduate departments from each of 10 to 15 fields representing a wide range of verbal vaquantitative emphasis (for example, engineering to English), with some fields of relatively mixed emphasis such as education and biology.

The records of examinees eligible for inclusion in the study (by enrollment, citizenship, language status, and data-availability criteria) were classified by reported undergradaute major field, and the fields were ordered in terms of the total number of designators. Within each field classification, examinees were distributed according to designated undergraduate school, and schools were ordered according to total number of designators without regard to field—that is, in terms of total volume of graduate-school bound, currently or recently enrolled students in the GRE pool.

The 20 most frequently designated fields are listed below, and those selected for the study are identified by asterisks:

sociology* mathematics* music electrical	computer science* other biosciences other social sciences physical education agriculture*
	mathematics*

English, history, sociology, and political science may be thought of as representing primarily verbal fields; chemistry, computer science, mathematics, electrical engineering, and economics were selected as representing primarily quantitative fields; and agriculture, biology, and education represent fields not clearly classifiable according to relative verbal and quantitative emphases.

Schools and departments we selected, within each of the 12 field classifications, by specifying certain minimum Ns, set after inspection of the data, to lead to inclusion of 20 or more samples from undergraduate schools contributing varied numbers of students to the general GRE examinee pool. Results of the selection process are indicated in Table 1. Pata on



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Distribution of Undergraduate Departmental Samples Included in the Study

By Jize and Field

Number of undergraduat - departmental samples by field Comp Eng-Socio-His-tory Polic Agri-d Biolg Mathe-c Elec_ Koucaz ALI lish size tetry Karin CEY tion fields 100+ 90-99 50-49 70-79 60-69 50-59 40-49 3 10 21 30-19 2 11 31 20-25 16 5 7 5 12 13 33 124 10-19 33 24 * 24 16 38 34 13 15 36 233 **410** 10 10 No. of 63 39 26 45 25 41 123 24 51 36 44 40 437 depte No. of 364 545 544 647 251 850 663 976 1318 1649 9375 etudente Male(2) 54.8 25.8 57.2 67.2 69.6 62.5 88.3 62.9 59.7 45.9 12.0 49.4 Minority(I) 11.0 13.9 29.2 18.8 14.6 17.9 11.0 21.6 15.4 7.3 14.9 9.0 14.1

Mote. An undergraduate departmental sample includes individuals naming a designated undergraduate major field and a designated undergraduate school who were taking the GRE General Test during 1981-82 as either (a) enrolled undergraduates or (b) nonenrolled bachelor's degree holders no more than two years beyond the bachelor's.

Minimum N = 15; Minimum N = 10; Chinimum N = 9; dhinimum N = 20

sex composition and minority representation in the sample, by field, are also shown.

As may be determined from Table 1, the study sample included 9,375 individuals from a total of 437 undergraduater departments in 149 different undergraduate institutions. In 8 of the 12 fields, the model number of undergraduate majors per department was between 10 and 19, and distributions of Ns per department were positively skewed around these small model values within each field. These conditions are quite similar to those encountered in graduate level validity studies.

GRE Item-Type Part Score Data

For each member of the study sample, operational GRE scal I verbal, quantitative, and analytical scores and corresponding item response data were available, based on one of six different forms of the GRE General Test that were used during 1981-82. Each form included the same total number of items, and the same number of items by type, as indicated below:

Variable	No. of
Verbal Test	(76)
Antonyms	22
Analogies	18
Sentence completions	14
Reading passages	22
Quantitative Test	(60)
Quantitative comparison	30
Regular mathematics	20
Data interpretation	10
Analytical Test	(50)
Analytical reasoning	38
Logical reasoning	12

Raw total scores (based on the 76 verbal, 60 quantitative, and 50 analytical test items) were computed for each member of the study sample taking each form of the test, and raw part scores were computed for each of the nine item types indicated above; in addition, a vocabulary score based on the 40 antonyms and analogies items and a reading comprehension score based on the 36 sentence completions and reading passage sets were computed for each individual. All raw scores were computed using the total number right scoring procedures introduced during 1981-82.

The part scores are of differing lengths, with corresponding differences in reliability. For example, based on internal analyses of two forms of the GRE General Test administered during 1981-82 (Wallmark, 1982a; 1982b),



typical levels of Peliability (estimated by Kuder-Richardson Formula 20) of the various GRE scores in general smaples of GRE examinees are approximately as follows:

Test	Typical form reliability						
Verbal Test (Total)	.90+ (76 items)						
Antonyms + Analogies +	•						
Sentence completions	.90 (54 items)*						
keading comprehension	.80+ (22 items)						
Quantitative Test (Total)	.90 (60 items)						
Quantitative comparison	.80+ (30 items)						
Regular mathematics	.75+ (20 items)						
Data interpretation	.60+ (10 items)						
Analytical Test (Total)	.85+ (50 items)						
Analytical reasoning	.80+ (38 items)						
Logical reasoning	.60+ (12 items)						

and a 36-item reading comprehension score would each have reliabilities exceeding .80 in samples such as those employed in the internal studies cited. Since the validity of a test is partially a fuention of its reliability, the differences in reliability should be kept in mind in evaluating the validity of the various part scores—that is, a shorter test of a given ability may be expected to have somewhat lower validity than a longer test of that ability, given a common external criterion. For purposes of this study, reliabilities approximating these noted above are assumed to obtain for the various measures.

Preliminary operations on the raw GRE total and part scores. In evaluating the predictive validity of operational GRE verbal, quantitative, and analytical scores, the fact that the scores are based on different forms of the test does not pose problems of score comparability across forms. Through a process of test equating, raw total scores ea ed on each new form of the GRE General Test are placed on the GRE scale by means of formulas that calibrate the scores to make them comparable with those on earlier forms, regardless of differences in the level of difficulty of the respective forms (for example, ETS, 1981).

Towever, equating procedures involve only the raw total scores on the live forms of the est—different sets of item types within a test are researily parallel in difficulty in a given form, and sets of items of an type are not necessarily parallel in level of difficulty across



^{*}In internal analyses, sentence completions are combined with analogies and antonyms for statistical evaluation.

forms. Thus, combining raw-score data across forms without formal equating introduces some elements of interpretive ambiguity into a validity study. The analysis could have been conducted using only data from a single test form (obviating interpretive complications) but this was not considered desirable because use of single-form data would have substantially restricted sample size and because there might be differences across forms and administrations in examinee mix with respect to variables such as sex, educational status at time of testing, selectivity level of undergraduate school attended, and the like—variables that could have some bearing on study outcomes.

Formal equating of the raw part scores was not feasible for this exploratory study. Without resolving questions regarding the relative difficulty of the respective item types within and/or across forms, it was decided to transform raw part and total scores to a common scale, by form, with full awareness of the attentuation in validity that might be associated with this procedure. In this regard, it was assumed that item types differ only randomly, within and across forms, with respect to parallelism. It was also assumed that attenuating effects due to lack of practicular sets of items. (See Table 14 and Appendix B for evidence bearing on these assumptions.)

Based on data for examinees taking each form of the GRE General Test without regard to their field of study, raw part and total score distributions were subjected to a z-scale transformation (mean = 0.0, standard deviation = 1.0)—that is, raw part and total scores were expressed as deviations from the respective form grand means in standard deviation units, using the means and standard deviations shown in Appendix A.*

It was reasoned that validity coefficients for the z-scaled part and total scores would be attenuated by any errors associated with lack of equating, while coefficients fdor the GRE scaled (converted, fully equated) total scores would not. It was assumed that comparison of validity coefficients for the total scaled (equated) scores with those for the z-scaled (uequated) total scores would indicate the overall effect on validity of combining unequated total (and part) scores across forms. It was assumed further that, for comparing the validity of total test scores with that of various part scores, the appropriate total scores would be the z-scaled transformations of the raw total scores (paralleling transformations of the respective part scores) rather than the converted GRE scaled total score.



^{*}Appendix A also provides data on the number of examinees taking each form by sex. These data indicate pronounced differences in Jex mix across forms and administrations; males constituted a majority of examinees taking forms administered in October, December, and February while females constituted a stronger majority of those taking forms administered in April and June. By inference, differences in major-field mix may also be present.

In summary, the test variables available for study following the operations described above were as follows:

```
V GRE scaled verbal score (equated across forms)
Q GRE scaled quantitative score (as for V)
A GRE scaled analytical score (as for V)
V* Standardized raw total verbal score (not equated by form)
Q* Standardized raw total quantitative score (as for V*)
A* Standardized raw total analytical score (as for V*)
```

Standardized raw item-type part scores:

```
(antonyme)
ANT
ANA
     (analogies)
SC
     (sentence completions)
RD
     (reading passages)
VO
     (vocabulary or ANT + ANA)
RC
     (reading comprehension or SC + RD)
OC
     (quantitative comparison)
RM
     (regular mathematics)
DI
     (data interpretation)
AR
     (enalytical reasoning)
LR
     (Togical reasoning)
```

Finally, one additional set of GRE "total scores" (designated V#, Q#, and A#, respectively) was included, namely, one in which the various item-type part scores were given equal weight. Given the z-scaled part scores, total /erbal, quantitative, and analytical scores defined by the sum of their respective parts were computed for each member of the study sample. In these total scores, item types are equally weighted since the standard deviations of the z-scaled scores are identical. If validity coefficients for V#, for example, should exceed those of, say, V or V* (in both of which the item-type subtests are weighted according to their I ngth), then it may be concluded that the current relative representation of the respective parts in the total score is not consistent with their relative contribution to prediction.

Study Procedures

As indicated earlier, scores on the study variables were available for 437 undergraduate departmental samiles, distributed among 12 major fields. In order to assess similarities and differences among the major-field classifications, without regard to department of undergraduate enrollment, profiles of means on the z-scaled item-type part scores were developed for the 12 major-field groups. Questions regarding the relationship of the respective test measures to the SR-UGPA criterion were explored using scores that were first standardized by department and then pooled across all departments within the respective fields of study.



Pooling rationale. Results of regression analyses in small samples are subject to substantial sampling fluctuation. By pooling data for several small samples from similar setting (for example, several undefgraduate chemistry departments), it is possible to obtain more reliable estimates of relationships than would be possible in single small samples. In pooling data across departments one useful approach has been to standardize the predictor and the criterion variables within each department before pooling—that is, to express scores on all variables as deviations from department means in department standard deviation units (see, for example, Wilson, 1979; 1982). For each departmental sample, the mean on each variable is zero and the standard deviation is unity.

Coefficients computed for pooled departmentally standardized variables, by field, may be thought of as approximating population values around which the coefficients for individual departments will vary due to selection—and sampling—related considerations (for example, restriction of range on predictors) as well as context—specific validity—related factors (for example, economics departments may differ in curricular emphasis on quantitative methods of analysis).

A majority of the varietion in observed validity coefficients in samples from similar settings tends to be accounted for more by statistical artifacts than by situation-specific validity-related factors. For example, it was found that about 70 percent of the variation across 726 validity studies in the correlation between Law School Admission Test scores and first-year law school grades was attributable to differences in sample standard deviations, estimated criterion reliability, and sample size (Linu, Harnisch, & Dunbar, 1981). Similar findings have been reported for employment settings (for example, Pearlman, Schmidt, & Hunter, 1980).

When analyses are based on pooled, departmentally standardized data within a given field of study, emphasis is on identifying the characteristic patterns of relationships between the respective GRE variables and the measure of academic performance under consideration.

Major-Field Differences in Average Performance on GRE Item-Type Subtests

Table 2 show, means on the GRE verbal, quantitative, and analytical item-type part scores and the respective total scores for examinees in the 12 major fields of study. For all except the converted (GRE scaled) verbal, quantitative, and analytical total score means, the means indicate the average deviation of the raw scores of examinees in a given field from the mean of all examinees in the study sample without regard to field, in all-examinee standard deviation units.

Thus, for example, undergraduate English majors were .622 standard deviations above the all examinee mean on the verbal test (STNRAW V* = .622), .376 standard deviations below the all examinee mean on mean quantitative ability (mean STNRAW Q* = -.376), and so on. Similar interpretations may be made for other means in the table.



Table 2
Means for Major Field Groups on Test Variables

	Verieble	fin- glich	His- tery	Roct- elegy	Po- lit- ical Sci	Chan- Istry	Com- put- ar ses	Meth	ilee-	Rev- nou- ion	Biel-	Agri- cul- ture	Edo- eo- Edon
											•		
	Converted Verbal	360,7	578.2	460.7	536,1	534.3	527.4	546.3	,' 511.2	, 333.1	537.2 4	is.3	438.3
	Conversed Quantitative	532.4	559.1	467.8	349.8	656.4	447.9	700.9	703,6	:23.e	965.2 5	50.2	466.6
	Converted Analytical	566.3	576.6	484.0	(40.6	451.7	625,9	634.1	614.1	602.0	557.3 5	43.5	. 407.4
	STIRLE-Y	. 622	.302	-4520	.157	. 124	.001	.247	.111	.313	.143	357	672
	STREET-Q*	-, 376	183	-,506	-,245	.591	, 794	,964	.976	, 300	.198	-,155	125
	man-1	623	.054	485	070	.343	.483	.547.	.300	.274	. 145	~,267	~ 566
												i	
	Anto-rue	.672	.543	383	172	.741	500	.197	.007	.233	.117	355	546
	Annin (San	. 396	.440	459	. 546	.006	.092	.228	.948	.240	. 21.5	~. P.S	556
	Sentence Completions	. 525	.401	416	-157	.119	.071	. 222	.141	.273	.110	-,234	~.583
	Inediag Passages	. 337	. 325	-,503	'ôtr	. ',67	, 242		.191	.254	.204	111	596
	Tocalesary	.699	,546	-,448	.176	.072	,847	. 872	,826	"OS4	.136	~375	463
	Booking Comprehentes	.431	. 365	518	.113	.162	. \$20	est.	.24	. 33.9	.183 .	237	-, 636
7	Commitative Comparison.	. 326	153	-,837	217	\$26.	.652	.924	.936	. 366	.172	163	01
	Regular Math	.417	-, 344	- ,807	291	.566	.696	.+77	,.51	. 335	-177	154	728
	hate laterprotetion	215	731	~ .56)	091	.392	. 46 's	.523	,594	. 370	.151	035	643
	Analytical Secretary	130	443	-,669	446	. 345	.531	.569	- 416	.198	. 155	139	-,479
	logical flatorning	.276	.284	404	.33	. 195	. ,155	.268	,163	.367	.038	286	,698
	y	(654)	(\$44)	(364)	1342	- (644)	(847	§ -{291]	} (45 8)	(6.2)	(1314)	(976)	(1549)

Note: Converted V, Q, and A are operational GRE-s saled scores, fully equated across all forms administered during the year.

STNRAW V, Q, and A are within-form standardizations of unequated crv total scores on the respective tests. Raw total scores were z-scaled by form using data for all examinees taking with 100% without regard to their fields of study.

#Except for the converted V, Q, and A sucres. all test scores were x-scaled by form of test taken, using data for all examinees taking a form without regard to field. Thus, the grand mean for all ex minees is 0.0 and the atandard deviation of such tust distribution is unity. The means reported indicate the deviation of the mean for a given major field group from the all-examinee mean in standard deviation units. Thus, for example, the mean of .672 for antonyme reported for English majors indicates that they were .672 standard deviations above the grand mean on this variable, on the average.



Figure 1 highlights differences among and within fields in performance on the item-type part scores. Profiles for majors in the four humanities and social sciences fields and in education (thought of as verbal fields) are shown together in the left portion of the figure; those for majors in the four math and science fields and economics (thought of as quantitative fields), and in biology and agriculture (thought of as fields with mixed or balanced quantitative and verbal emphases) are shown in the right portion of the figure.

Within-field differences in level of performance on the item-type part scores are of particular interest. For example, majors in the verbal fields typically performed better on the vocabulary items (ANT and ANA) than on the reading comprehension items (SC and RD); they performed better on data interpretation (DI) items than on quantitative comparisons (QC) and regular mathematics (RM) items; and, with the exception of majors in education, they performed at a sharply higher level on logical reasoning (LR) than on analytical reasoning (AR) items.

Majors in chemistry, mathematics, engineering, and computer science tended to exhibit an opposite pattern, with higher performance on reading comprehension items than on vocabulary items, higher performance on quantitative comparisons and regular mathematics than on data interpretation items, and much better performance on analytical reasoning than on logical reasoning items. Mathematics majors differed from the others in this cluster primarily by performing considerably less well on reading passages (RD) items than on sentence completion (SC) items.

Verbal part-score profiles for majors in economics, biology, and agriculture tended to parallel those for the math and science fields (better on reading comprehension than vocabulary); on the quantitative part scores, their profiles do not exhibit the extreme contrast between quantitative comparisons, regular mathematics, and data interpretation items characteristic of profiles for the math and science majors. With respect to items in the analytical test, agriculture and biology majors, like math and science majors, performed better on analytical than logical reasoning items, but economics majors, like the verbal majors, had a higher logical reasoning than analytical reasoning mean.

Another way of assessing variability in major-field performance on item-type subtests within the respective ability measures is to examine (a) the relative standing of the several major field groups in terms of means on the subtests within a test and (b) the absolute differences in means for various pairs of subtests. For example, for two parallel tests a high degree of consistency in the ranking of field means and relatively small absolute differences in corresponding z-scaled means would be expected; a lower degree of consistency in field ranks combined with larger differences in z-scaled means, on the other hand, would be expected for tests measuring different abilities.

Table 3 shows for pairs of subtests within the respective tests, (a) whether the ranks of the 12 major fields were identical or different and the absolute difference in the ranks when differences were present and (b) the absolute difference in z-score means. The absence of an entry in the rank



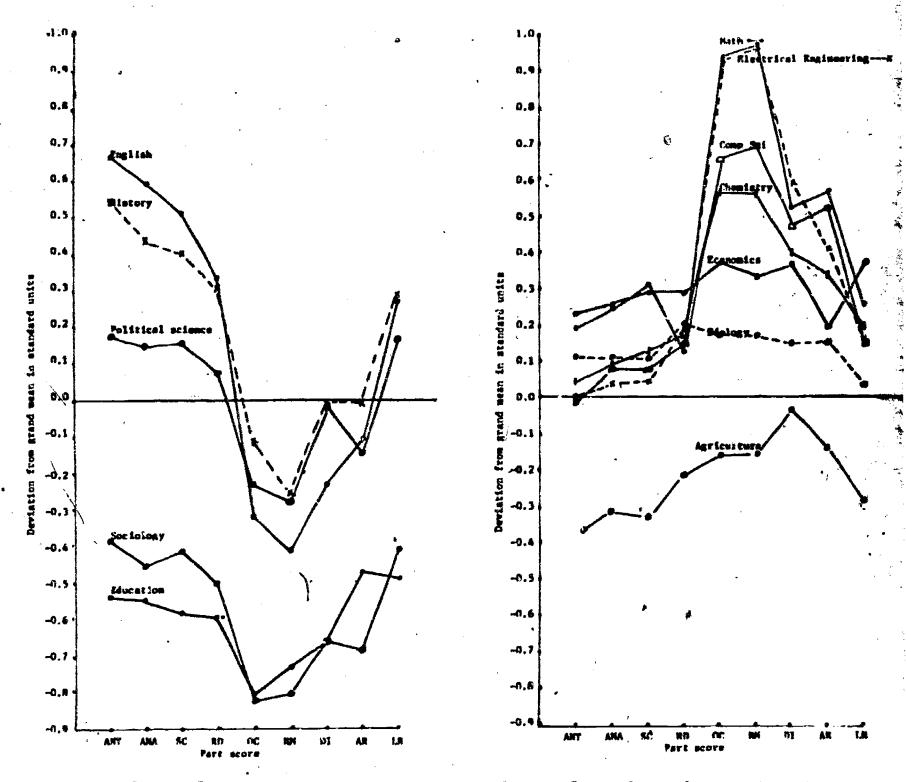


Figure 1. Profiles of mean scores on GRE item-type subtests for undergraduate majors in the fields selected for study

Source: Table 2



Table 3

*Observed Absolute Differences in the Ranks of Heans of 12 Major Field Groups on Pairs of Item-type Subtests, By Test, and

Associated Absolute Differences in I-score Heans

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Fairs of part scores by test Field Verbal Quantitative Analytical M-DI OC-DI W-RC Renk Hean Rank Nean Mean Rank Mean Rack Meso Rank diff. diff. Diff. Diff. Diff. Diff. Mil. Diff. piff. Diff. .248 .000 .092 5.0 .406 1.0 1.0 .189 English .325 5.0 History .161 1.0 .096 .122 1.0 .218 .292 .024 0,72 Sociology .056 .174 .146 1.0 .309 .074 .128 .200 3.5 4.0 .063 Pol. Sci. 162 .095 .002 .164 1.0 .150 Chemistry .044 .229 .077 .185 6.0 .376 Comp. Sci. .001 .053 .401 1.0 .454 3.0 .309 Math. 1.0 .332 .367 3.5 .253 4.0 - .162 1.0 .035 1.0 Elec. Eng. .054 .034 .001 ...035 4.0 .169 Economics .088 .128 .047 1.0 .001 .199 Agriculture Biology **~057** .034 .022 .026 3.0 .097 .028 .045 .138 .064 1.0 - .019 Education

Note: No entry in the rank difference column indicates that the means for the major field group on the designated pair of subtests had identical rank. If there was any discrepancy in rank, the entry indicates the absolute difference between the ranks. Entries in the mean difference column indicate differences between means in standard units (absolute differences).

(,034)

(.128)

(.175)

(.300)

Source of data: Table 2

(.067)

Hedian (diff

in mean)

Table 3.1

Intercorrelations of the Ranks of Means on Item-type Part Scores

for 12 Major Field Graups

	VO	RC	QC	TIK	DT *	AR	12
VO		.858	.182	.161	.187	.315	.897
RC			.406	.357	.411	.510	. 697
QC				.986	.999	.958	.403
RM				-	.984	.958	.362
DI						.956	.406
AR							.481
i.E						`,	-

Note: Entries are rank correlation coefficients (rho).

difference column indicates that the ranks of the major field gorup on the designated pair of tests were identical. For additional perspective, Table 3.1 shows rank correlations (rho) of subtest means for the 12 major fields, across as well as within tests.

- o Considering the two verbal subtests, VO and RC, for 10 of the 12 fields, ranks of z-scaled means were identical and the median absolute difference in z-scaled means was relatively small (.067). The rank correlation (rho) for the 12 field means on these two variables was .888 (Table 3.1).
- o For each of the three pairs of quantitative subtests, there were relatively minor discrepancies in rank order, with no shift of more than one rank. For QC and RM, the median absolute difference in means was quite small (.034); however, median differences in field means were greater for both QC and DI (.128) and RM and DI (.175). The average rank correlation of field means (Table 3.1) for the three pairs of quantitative subtests approached .99.
- o For the two analytical ability subtests, AR and LR, some shifts in ranking were found for every field, the median absolute difference in z-scaled means (.300) was higher than that for subtests within the verbal and quantitative ability measures, and the rank correlation of field means on AR and LR (rho = .481) was considerably lower than that for the other pairs of subtests.

The findings regarding field means indicate the differential development within individuals, associated with field of concentration, of the skills and abilities being measured by different item types within the respective tests. On balance, the evidence reviewed in this section suggests that the item-type part scores are not simply different methods of measuring their respective constructs but that they may represent distinguishable components of underlying general abilities with the potential for independent measurement utility.

In this connection it is important to note (a) that the degree of consistency in major field performance differentials is greater for subtests within the verbal and quantitative ability measures than for the two analytical ability subtests, (b) that the field ranks on analytical reasoning items correspond closely with ranks on all three quantitative item types (average rho of approximately .960), and (c) that field ranks on logical reasoning items correspond closely with ranks on the two verbal subtests (rho = .8967 for both LR-VO and LR-RC). Generally speaking, the rank correlations in Table 3.1 indicate that, insofar as major field performance differences are concerned, the information conveyed by the analytical reasoning and the quantitative subtests is similar and that conveyed by the logical reasoning subtest and verbal subtests is also similar.

Exploratory Evaluation of Part-Score Validity

The analyses involving part-scores on the verbal measure were guided by



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several a priori working hypotheses, based on the College Board findings cited at the outset:

- 1. The GRE reading comprehension (RC) subtest (based on sentence completions and reading comprehension sets) should be more closely related to SR-UGPA than the GRE vo_abulary (VO) subtest (based on antonyms and analogies).
- 2. The 36-item RC subtest should be comparable in validity to the total GRE verbal test, including the 40 VO items.
- 3. The mainly correlation of the RC, Q^* , A^* battery with SR-UGPA should be comparable to that of the V^* , Q^* , A^* battery.
- ccasional suppression of VO, but not RC, variance may be expected in or cosites including RC, VO, and other GRE variables.

In the absence of comparable working hypotheses regarding the quantitative and analytical part scores, evaluation of observed relationships for these item types was guided by interest in (a) the relative contribution of the respective item-type part scores within each test to prediction of SR-UGPA, (b) the comparative validity of total test scores and the component part scores, and (c) evidence suggesting the possibility that separately scored item-type subtests might provide a basis for improved assessment.

The Verbal Test Part-Score Analysis

Table 4 shows pooled within-department correlations between SR-UGPA and (a) VO and RC scores, (b) various verbal total scores, and (c) a best-weighted combination of VO and RC scores, by field, and for all fields combined. Validity coefficients for V* (the raw unequated total verbal score, z-scaled by test form) were slightly lower than those for V (the converted, GRE-scaled operational verbal score). This outcome is expected because V* total scores, like the respective part scores, were not equated across forms. In comparing part- and total-score validity, V* is judged to be the more appropriate total, under the assumption that attenuating effects associated with lack of equating across test forms are comparable fdor V* and the respective part scores. . Coefficients for V# (a total defined as the sum of equally weighted scores on analogies, antonyms, sentence completions, and reading sets) and V* are assumed to be comparably attenuated due to errors associated with lack of equating across forms. This same line of reasoning is applicable, of course, to later consideration of data on the quantitative and analytical measures.

The validity coefficients for the verbal measure varied by field generally in accordance with the expectation of higher validity in the more verbal fields than in the more quantitative fields. This was true without regard to the particular verbal measure under consideration. However, for



Pooled Within-Department Correlations of Selected Verbal

Para and Total Scores with SR-UGPA, by Field

Field .	(N)	Verbal	•	\$,	Verbal scor	total		· Differen	ice in v	aliditý
		· VO	RC		v* r	V∜ r	V	VO,RC	RC vs VO	V* vs RC	VO.RC Vs V*
English	(884)	347	377		395	395	399	399°	030	018	004
History	(584)	322	354	•	366	370	377	375	032	012	009
Sociology	(364)	342	396		407	384	418	417	054	011	010
Polit Sci	(545)	283	376	:	362	364	364	380	093	-014	018
			•								1
Chemistry	(644)	226	217		243	242	249	248	-016	026 '	- 005
Computer Sci	(647)	238	213	:	246	209	245	251	-025	033	005
Mathematics	(251)	248	312	;	292	296	. 321	314	064	-020	022
Elec Engin	(850)	140	249	•	21.1	201	223	253	109	-038	042
Economics	(663)	323	391	•	391	390	404	403	068	000	012
Biology	(1318)	228	288		286	269	302	297	060	-002	011
Agriculture	(976)	214	215	;	239	236	260	239	001	024	000
Education	(1649)	296	313	:	333	326	332	335	007	010	002
All Fields	(9375)	263	301		309	300	318	315	038	008	006

Note. V* is the raw total verbal score, z-scaled by form.

V# is an equally weighted sum of four verbal part scores.

V is the converted GRE scaled verbal score, equated across forms.

VO,RC is a best weighted composite of the designated part scores.

Entries are correlation coefficients without decimals.

economics, among the more quantitative fields, the verbal test had validity coefficients comparable to the coefficients for the English, history, sociology, and political science samples.

With respect to patterns of verbal part-score validity, the findings in Table 4 are generally consistent with the basic working hypotheses outlined above.

- o Considering first the all-fields coefficients (equivalent to weighted averages of coefficients for 437 departments without regard to field), the validity for RC is greater than that for VO (coefficients differ by .038, as indicated in the RC vs VO difference column), and this was true for 10 of the 12 fields. For chemistry and computer science departments, the mean VO coefficient was higher than the mean RC coefficient, but the mean difference was less than the average for all departments in absolute magnitude.
- o RC slone was about as valid as V* including the VO items. Considering data for all departments, without regard to field, coefficients were .301 and .309, respectively. RC was actually slightly more valid than V* in several fields.
- o However, a best-weighted composite of VO and RC did not yield much better prediction than the V* total score, similar to the results observed with SAT vocabulary and reading comprehension when they were similarly treated. Largest differences in validity between the VO,RC composite and V* occurred in three of the four fields in which RC was more valid than V*, and in which differences in validity between RC and VO were greatest, namely, electrical engineering, mathematics, and political science.

The data in Table 5 lend support to the working hypothesis that the multiple correlation of an RC, Q^+ , A^+ composite with SR-UGPA should be comparable to that of a V^+ , Q^+ , A^+ composite.

o For all departments, without regard to field, the coefficient for RC,Q*,A* was only .002 points less than that for V*,Q*,A*, and .005 points less than that for VO,RC,Q*,A*—that is, when VO was added to the RC,Q*,A* battery there was little increase in the multiple correlation. There were no notable exceptions to this general finding by field.

Table 6 provides evidence regarding the relative weighting of two sets of verbal part scores, namely, VO and RC (Set 1), and the four basic verbal item types, namely, analogies antonyms, sentence completions, and reading comprehension sets (Set 2), when included in a battery with Q* and A*.

o The data in Set 2 indicate, among other things, (a) that, over all departments, the relative weighting of sentence completions and reading items (components of the RC score) was approximately equal, (b) that one of these two RC item-types was the highest of the four verbal item types in all fields but one (agriculture), but (c) that the relative weighting of the SC and RD items, when they were allowed to compete independently, varied across fields without regard to their verbal or quantitative emphasis.



Table 5

Multiple Correlation with UGPA of Quantitative,

Analytical, and Selected Verbal Scores, by Field

-		" Multipl	e corre	lation	Di	fferenc	e .
	*	(1)	(2) VO.RC.	(3)	(3-1)	(3-2)	(2-1)
•		.RC,Q*,A*	Q*,A*	V*,Q*,A*		•	
English	(884)	384	403	402	018	-001	019
History	(584)	362 ⁸	382 ⁸	374 ⁸	012	-008	020
Sociology	(364)	436	447	442	006	-005	011
Political Sci	(545)	419 ^a	420 ^a	410 ^a	-0Q9	-010	001
Chemistry	(644)	372	375	374	002	-001	003
Computer Sci	(647)	365	374	371	006	-003	009
Mathematics	(251-)	412 ^a	412 ^{ab}	395 ^a	-017	-017	000
Elec Engin	(850)	393	406 ^a	386	-007	-020	013
Economics	(663)	452	458	455	003	-003	006
				•	;		
Biology	(1318)	352	354	350	-002	-004	002
Agriculture	(976)	.299	306	306	007	90 <mark>0</mark>	007
Education	(1649)	356	366	366	010	-001_	001
All Fields	·(9375)	 361	366	363	002	-003.	1005

Note: Entries are multiple correlation coefficients or differences between designated coefficients without decimals.

VO = ANT + ANA = Vocabulary

RC = SC + RD = Reading Comprehension

 V^* , Q^* and A^* are raw total scores on the respective tests, z-scaled by form.

[&]quot;A* variance is suppressed in this composite

byO variance is suppressed in this composite

Table 6

Relative Weighting of Two Sets of Verbal Part Scores, Quantitative and Analytical Scores in Composites for Predicting UGPA, by Field

Field	(N)		Set 1 . Bets weights (R)				Set 2						
• 14,10	(N)	vo				(R)	4 8 7 8	4 6 100	Beta	weight.		,	(R)
		VU	RC	Q*	A*		ANA	ANT	SC	- RD	Q*	A*	
English	(884)	166	230	061	017	(403)	104	077	168	168	063	010	- (411)
History	(584)	159	<u>238</u>	089	-045	(382)	037	101	241	072	078	-039	(398)
Sociol	(364)	126	215	035	171	(447)	142	031	056	168	034	169	(451)
Pol Sci	(545)	038	256	232	-054	(420)	093	-045	108	180	<u>230</u>	-060	(425)
Chem-	(644)	064	031	279	072	(375)ء	020	054	, 007	030	277	072	(376)
CS	(647)	106	011	274	064	(374)	059	063	-06^	065	274	059	(376)
Math	(251)	-012	237	300	-034	(412)	-128	077	209	098	296	-016	(431)
Elec E	(850)	- <u>147</u>	179	320	065	(406)	-089	-061	068	123	320	066	(406)
Econ	(663)	<u>099</u>	196	144	143	(458)	040	066	102	129	143	145	(458)
				•			•	•					
Biol	(1318)	050	158	187	054	(354)	028	036	033	142	189	052	(357)
Agric	(976)	<u>085</u>	038	178	076	(306)	026	065	050	009	179	073	(309)
Educ	(1649)	. 122	<u>131</u>	148	048	(366)	099	040	103	053	142	048	(371)
All Fields	(9375)	<u>079</u>	145	185	047	(366)	044	044	079	<u>090</u>	185	047	(366)

Note. Entries are standard partial regression (beta) weights and multiple correlation coefficients without decimals.

VO = ANT + ANA = Vocabulary: RC = SC + RD = Reading Comprehension

Q* and A* are raw total test scores, z-scaled by form:

Negative weights reflect suppression of variance; zero-order coefficients are positive. Underscored weights are estimated to be significant at the .05 level.

- o Suppressor effects, indicated by negative regression weights for predictors that are positively correlated with a criterion, were present for 'VO and/or VO component item types in analyses for mathematics, engineering, and political science departments, consistent with the hypothesis of occasional suppressor effects for vocabulary items; in one analysis (computer science departments), the sentence completion subtest was negatively weighted, contrary to hypothesis.
- o The Set I and Set 2 multiple correlations are (dentical in the analysis over all departments and are almost so in the 1. spective field analyses.

The analyses reviewed above indicate differences in the criterion-related validity of the VO and RC subtests favoring the RC subtest, which appears to be carrying most of the predictive salidity load in the total verbal score when the criterion is SR-UGPA.

The Quantitative Test Part-Score Analysis

Table 7 provides data on the relationship of the three quantitative item-type part scores to SR-UGPA. The correlations of three quantitative total scores, namely, Q*, Q*, and Q, with the same criterion are also shown. As expected, the validity coefficients for the various quantitative total scores are higher for the math and science and economics departments than for the other, less quantitative fields; however the higher validity of quantitative scores for political science departments than for other verbal departments was not expected.

In the absence of an a priori basis for expecting particular patterns of differential validity for the respective item types, perhaps the most relevant general consideration to be kept in mind is that the three quantitative subtests differ in length. QC includes 30 quantitative comparison items, RM includes 20 regular mathematics items, and DI includes 10 data interpretation items. Thus, we would expect validity coefficients to vary with test length if the three item-types are actually homogeneous with respect to the abilities they tap.

- o For all departments, the validity patterns for QC, RM, and DI followed the variation-according-to-length hypothesis, and this was true for several of the ffeld analyses as well. However, there were exceptions. For example, RM validities were somewhat higher than those for QC in several fields, most notably so in mathematics; DI validities comparable to those for QC were obtained in analyses for agriculture, English, and sociology (which are among the fields in which students performed better on the DI subtest than on the QC and RM subtests—see Figure 1).
- o A composite of the separately weighted part scores did not result in better prediction than that provided by Q*, based on the analysis over all departments. Only in the analysis for mathematics departments, in which the regular mathematics items had uniquely high validity, was there a notable exception to the foregoing generalization. The content



Table 7

Pooled Within-Department Correlations of Quantitative Part and Various Total Scores with UCFA, by Field

Field	(N)	Qua	ntitati score	ive part	Quantitative total scores					
		QC	, RM	DI	Q*	Q#	·Q	QC,RM, DÍ		
	,	r	r	r	r	·r	¥	R		
English	(884)	209	209	192	238	241	246	245		
History	(584)	16	203	126	212	205	225	224		
Sociology	(364)	226	259	216	285	293	310	286		
Polit Sci	(545)	353	269	216	353	335	362	362		
Chemistry	(644)	330	305	203	358	340	371	366		
Computer Sci	(647)	285	293	258	350	349	356	350		
Mathematics	(251)	294	366	210	356	340	378	382		
Elec Engin	(850)	346	290	212	378	356	397	380		
Economics	(663)	307	283	212	348	339	358	348		
Biology	(1318)	268	246	182	296	287	310	298		
Agriculture	(976)	216	234	217	276	280	306	278		
Educ	(1649)	285	243	193	302	292	302	304		
All Fields	(9375)	274	257	201	308	300	320	308		

Note. Q* is the raw total quantitative score, z-scaled by form.

Q# is an equally weighted sum of quantitative part scores.

Q is the converted quantitative score, equated across forms.

Entries are correlation coefficients without decimals.

of the regular mathematics items may overlap more with the content of the major field for mathematics majors than for majors in the other fields. If so, this would help to explain the strong predictive validity of these items and would be consistent with findings of previous research indicating characteristically higher validity for the GRE Subject (Advanced) Tests then for the General (Aptitude) Test (see, for example, Willingham, 1974; Wilson, 1979; 1982).

Table 8 provides insight into the relative weighting of QC, RM, and DI when the three part scores were included in a battery with A* and V*. The difference in multiple correlation between the QC, RM, DI, A*, V* composite and the Q*, A*, V* composite is also shown. The predictive load, relative to the SR-UCPA criterion, in the quantitative test is being borne primarily by the QC and RM items, judging from the findings in Table 8.

o DI contributed only slightly to prediction, generally, and attained statistical significance only in the analyses for computer science and agriculture departments; suppression effects were-found for DI in analyses for two verbal fields (history and political science) and mathematics. In a stepwin- regression program, QC, RM, and DI were entered as a set followed quentially by the introduction of A*, then V*. In the three analyses showing DI suppression (and in all other analyses), the weight for DI was positive in the initial quantitative set. The DI weight became negative only after the introduction of the final variable (V*) in analyses for history and political science, but after the introduction of A* in the mathematics analysis.

o Separate treatment of QC, RM, and DI part scores in a battery with A^* and V^* did not lead to better prediction than that provided by Q^*, A^*, V^*

(see difference column in Table 8).

The Analytical Test Part-Score Analysis

The analytical ability measure introduced in October 1981 is a revised version of the analytical measure introduced when the GRE General Test was restructured in 1977. There is empirical evidence regarding the validity of the October 1977 analytical measure for predicting graduate school performance (for example, Wilson, 1982), but evidence regarding the October 1981 version is more limited. Evidence of positive relationships between SR-UGPA and analytical reasoning and logical reasoning items, respectively, was reported by Wild, Swinton, and Wallmark (1982) in studies leading to the revision of the 1977 measure. In those studies, logical reasoning items were found to be more closely related to SR-UGPA than analytical reasoning items in samples that were not differentiated with respect to field.

The analyses reported in this section provide evidence regarding the relationship of the various analytical ability total scores (A*, A*, and A) and the component analytical ability item types, namely, analytical reasoning (AR) and logical reasoning (LR), to SR-UGPA in samples classified



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Table 8

Relative Weighting of Quantitative Item-Type Part Scores in a Composite with A* and V*

Field	(N)			QC,RM,DI	Increase			
		QC	RM	DI	A*	V*	A*,V* (R)	over Q*,A*,V*
English	(884)	-004	045	054	018	<u>355</u>	404	003
History	(584)	053	096	-033	-040	347	380	Ò06
Sociology	(364)	-037	058	013	182	301	444	002
Polit Sci	(545)	223	050	-007	-041	253	416	006
Chemistry	(644)	184	132	028	074	077	381	007
Computer Sci	(647)	116	126	100	061	101	370	000
Mathematics	(251)	094	265	-094	-021	191	418	023
Elec Engin	(850)	219	124	047	082	030	389	003
Economics	(663)	- 091	068	002	159	<u>260</u>	454	000
Biology	(1318)	<u> 126</u>	089	008	<u>063</u>	180	352	002
Agriculture	(976)	042	095	087	073	114	308	002
V domant dans	(1(10)		0.1.0	•				,
Education	(1649)	. 117	049	002	052	<u>226</u>	368	002
All fields	(9375)	<u>107</u>	089	025	<u>053</u>	197	364	001

Note. Entries are standard partial regression (beta) weights or multiple correlation coefficients without decimals.

Underscored weights are estimated to be statistically significant (p< .05).

by field of study. In evaluating the observed correlations, in Table 9, it is important to keep in mind that total scores on the 50-item analytical measure are more heavily influenced by performance on the 38 analytical reasoning items than by performance on the 12 logical reasoning items.

Generally speaking, typical validity coefficients for the various analytical total scores tend to be somewhat higher in the primarily quantitative fields (except mathematics) than in the verbal fields (except sociology). However, the AR and LR subtests do not have similar patterns of validity coefficients across verbal and quantitative fields.

In this regard, perhaps the most striking aspect of the part-score validity data in Table 9 is the strong contribution to prediction of Sk-UGPA, relative to that of the 38-item AR subtest, of the LR subtest based on only 12 logical resoning items.

- o For all departments, the validity of the LR subtest was .225 as compared to .229 for the longer AR subtest.
- o In seven analyses, the validity coefficient for LR was approximately equal to or greater than that for AR.
- o In three analyses (for history, political science, and education departments), the LR subtest validity coefficient was greater than that for the A* total (which included the AR items).
- o AR validities tended to be somewhat higher for the basically quantitative fields than for the basically verbal fields; for LR, validity coefficients tended to show an opposite pattern.

The relative weighting of AR and LR in an independently computed composite and their weighting in a composite with V^* and Q^* are shown in Table 10.

- o When AR and LR were treated as predictors, AR weights were somewhat higher than those for LR in composites for the chemistry, computer science, mathematics, electrical engineering, biology, and agriculture analyses.
- o LR weights were somewhat higher than AR weights in analyses for history and political science (among the more verbal fields), for economics alone among the more quantitative fields, and for education.

Although, when considered jointly as an independent battery, weights for both AR and LR reached the .05 level of statistical significance in most of the analyses, neither AR nor LR made a consistent, substantial contribution to prediction when treated as elements in a battery that included the V* and Q* total scores (cf. results in Table 6 for verbal subtests combined with A* and Q* and in Table 8 for quantitative subtests combined with V* and A*).

o Only the bets weight for LR was significant in the overall departmental analysis, and its contribution to prediction was relatively slight (bets * .058 as compared to approximately .190 and .185 for V* and Q*).



Table 9

Pooled Within-Department Correlations of Analytical Part
Scores and Various Total Scores with UGPA, By Field

Field	(N)	_	ical part ores	A	Analytical total scores					
		AR	. LR	A*	AV	A	AR, LR			
		r	r	r	r	r	R			
English	(884)	202	200	236	248	239	243			
distory	(584)	146	239	195	233	205	249			
Sociology	(364)	314	300	360	376	375	372			
Polit Sci	(545)	162	269	229	272	232	278			
Chemistry	(644)	255	175 ·	. 27 5	262	282	270			
Computer Sci	(647)	224	221	259	267	256	266			
Mathematics	(251)	218	214	239	250	251	263			
Elec Engin	(857)	264	199	282	273	292	287			
Economics	(663)	301	335	358	386	361	388			
Biology	(1318)	224	179	244	242	251	250			
Agriculture	(976)	220	184	240	238	255	246			
Education	(1649)	234	256	242	296	279	297			
All Fields	(9375)	229	225	264	274	270	275			

Note. A* is the raw total analytical score, z-scaled by form.

A# is an equally weighted sum of the analytical part scores.

A is the converted analytical score, equated across forms.

AR,LR is a best weighted composite of the designated part scores.

Entries are correlation coefficients without decimals.

Table 10

Relative Contribution of Ah and LR to Prediction of SR=UGPA in an Independent

Composite and in a Composite Including Q* and A*

Field	(N)	Beta :	weights	Multiple correlation		Multiple correlation			
	,	AR	LR	(R)	AR	LR	Q*	V*	(R)
English	(8 5.	148 '	145	(243)	012	013	070	353	(401)
History	(584)	074	214	(249)	-063	071	097	<u>319</u>	(381)
Sociology	(364)	241	214	(372)	133	107	031	284	(444)
Polit Sci	(545)	078	242	(278)	-102	088	248	231	(422)
Chemistry	(644)	221	094	(270)	069	-003	283	089	. (374)
Computer Sci	(647)	160	157	(266)	-002	092	289	083	(376)
Mathematics	(251)	163	158	(263)	-010	024	288	175	(395)`
Elec Engin	(850)	221	119	(287)	065	042	314	024	(387)
Sconomics	(663)	209	261	(388)	094	· <u>127</u>	147	230	(461)
Biology	(1318)	185	116	(250)	044	039	191	<u>173</u>	(351)
Agriculture	(976)	176	119	(246)	054	047	178	103	(307)
Education	(1649)	<u>161</u>	197	(297)	008	080	154	206	(370)
All Fields	(9375)	170	164	(275)	022	<u>058</u>	190	185	(365)

Note. Decimal points have been omitted from all coefficients. Underscoring indicates estimated statistical significance at the p \angle .05 level.

- o Suppression effects were found for AR in four departmental analyses and for LR in one; in the samples involved, AR or LR criterion-related variance was more than sufficiently represented in the verbal and/or quantitative total scores.
- o Weights for both AR and LR were statistically significant in only two field analyses (sociology and sconomics) and LR was significant in a third (education).

The data in Table 10 suggest that the analytical test, as currently defined by the 38 AR and 12 LR items, is not providing very much unique, SR-UGPA-related information. This conclusion is reinforced by the data in Table 11, which permit comparison of multiple correlations with SR-UGPA of V*Q* only and those yielded by adding A* and AR and LR, respectively. Increments in R due to adding analytical test scores to V* and Q* typically were quite small. In evaluating this finding, it is useful to know that V*Q* alone yielded a higher multiple correlation with SR-UGPA than either A*V* or A*Q* in 9 of the 12 field analyses and in the total sample.

Understanding of these findings is advanced by reference to Table 12 and Table 12.1. In Table 12 it may be seen that LR is more closely related to a verbal subtest (RC) than to LR. From Table 12.1 it may be determined that the average within-test intercorrelations of verbal subtests (.503) and quantitative subtests (.476) are greater than that observed for the two analytical ability subtests (.360); moreover, the correlation of LR with three of the four verbal subtests is higher than that of AR with these subtests while the correlation of AR with each quantitative subtest is higher than that of LR with these subtests. Intercorrelations corrected for errors of measurement shown in Table 12.1 (below the diagonal) lead to similar conclusions. In essence, AR items tend to have more in common with quantitative items than with LR items, while LR items have more common variance with verbal items than with AR items.

Werbal, Quantitative, and Analytical Part Scores as a Battery

Table 13 shows major findings of an analysis of the regression of SR-UGPA on seven item-type part scores, namely, VO, RC, QC, RM, DI, AR, and LR. Standard partial regression (beta) weights are shown for variables selected by stepwise regression as contributing at least .001 to R-squared.

- o The consistent significant contribution to prediction of the i and/or VO subtests is noteworthy; both are significant in four analyses, RC only is significant in five, and VO only in three (though acting as a suppressor in one).
- o The part score that appears to be contributing least to the battery is data interpretation (DI). However, the score for this subtest met the statistical significance criterion in the computer science and agriculture analyses.



Table 11

Incremental Contribution of the Analytical Measure (A*) in

Part-Score and Total-Score Form to Prediction of SR-UGPA

After Taking V* and Q* into Account, by Field

74.54	4	Ann Control	•					
Field	(N)	-	osite predi		Differen	Difference in R		
• . • .	•	(1) V*,Q* (R)	(2) V*,Q*, A*. (R)	(3) V*,Q* AR,LR (R)	(2-1)	(3–1)		
English	(1884)	401	401	401	000	000		
History .	_ (584)	374	374 ^c	381 ^a	000	007		
Sociology	(*364)	420	442	444	020	024		
Polit Sci	(545)	408	,410°	422 ⁸	002	014		
Chemistry	(644)	370	374	. 374 ^b	004	004		
Computer Sci	(647)	368	371	376 ⁸	003	008		
Mathematics	(251)	395	395 ^c	395 ^a	ροο	000		
Elec Engin	(850)	381	. 386	3 ა7	005	006		
Economics	(663)	439	455	461	016	022		
Biology	(1318)	346	*350	351	004	005		
Agriculture	(976)	300	306	307	006	007		
Education /	(1649)	364	366	370	002	006		
All Fields	(9375)	361	363	365	002	004		

Note. Entries are correlation coefficients without decimals.

^aAR negatively weighted

bLR hegatively weighted

cA* negatively weighted |

Table 12
Intercorrelations of Analytical Test Part Scores, and their Correlations with Selected Verbal and Quantitative Test Part Scores, by Field

Field	(N)	AR score vs	AR sc	ore vs	LR score vs		
		LE score	RC	QC	RC	QC	
٠		r	r	r .	r	r	
English	(884)	373	444	501	449	326	
History	(584)	337	404	497	472	321	
Sociology	(364)	341	401	502	442	276	
Polit Sci	(545)	352	436	476	468	404	
	•				•	,	
Chemistry	(644)	369	415	436	467	295 ີ	
Computer Sci	(647)	404	416	422	483	223	
Mathematics	(251)	346.	382	437	512	290	
Elec Engin	(850)	36 3	407	451	484	346	
Economics	(663)	358	387	461	508	366	
					:	•	
Biology	(1318)	_ 33,6	409	408	384	238	
Agriculture	(976)	368	451	478	488	321	
Education	(1649)	366	469	<u>557</u>	<u>507</u>	, 372	
All Fields	(9375)	360	429	475	469	318	

Note: Entries are correlation coefficients without decimals.

The higher coefficient in a given comparison is underscored.

Table 12.1

Pooled Within-Department Intercorrelations of
Item-Type Part Scores: Total Sample

	ANT	ANA	sc	RD	^o QC	<u>P</u> M	. DI	AR	LR
ANT L		528	493	486	277	266	233	282	356
ANA	843		509	486	337	290	260'	335	371
SC	711	784		519	336	274	267	332	394
RD	608 -	650	749	****	360	322	316	408	426
QC	433	372	485	450		548	440	475	318
RM	343	400	378	415	707		440	487	310
DI	336	359	388	456	635	651	_£	415	264
AR	441	441	407	510	594	628	600		360
LR	514	549	593	615	459	452	440	519	

Note: Values above the diagonal are observed correlations; those below are corrected for errors of measurement by use of the formula Tab/TaaTbb; reliabilities are estimated roughly.

Entries are correlation coefficients without decimals.



Table 13 . Beta Weights for Subsets of Item-Type Part Scores Selected by Stepwise Regression According to a Contribution to \mathbb{R}^2 Criterion, By Field

•		•	Part-	score	beta	weight	:s		Selected	V*,Q*,A*	N
Field	1	VO	RC	QC	RM	DI	AR	LR	Set (R)	(R)	
English		17	23		. 04	05			406	402 ^a	(884)
History		$\frac{17}{15}$	22	04	09		-08	06	392	376 ⁴	(584)
Sociology		12	21	•	06		11	10	451	442ac	(364)
Political Science			26	22	04		-10	80	437	410 ab	(545)
ALL VERBAL		12	23 22 21 26 23	05	<u>06</u>			05	(404)	(396 ^{ab})	(2377)
Chemistry		08		19		•	08	<i>'</i>	381	374 ^b	(644)
Computer Science		08 09		13	13	10		09	377	377 35	(647)
Mathematics		-		19 13 08	14 13 26 13 07	-			434	395ab	(251)
Electrical Engin	غد	-14	19	22	13	04	06		409	2060	(850)
Economics	₹	$\frac{-14}{09}$	18	$\frac{22}{08}$	07			12	463	, crade	(663)
ALL QUANTITATIVE			$\frac{22}{19}$ $\frac{18}{12}$	<u>15</u>	13	04	09 05	<u>12</u> <u>06</u>	(391)	387abc	(3055)
Biology		05	17	13	09 .		05		- 356	350 ab	(131s)
Agriculture				13 05	<u> </u>	09	06	05	309	30820	(.976)
ALL BALANCED		10 06	11	09	09 10 09	· <u>09</u> · 04	05	04	(332)	(330 ^{ab} c)	(2294)
- '				<u> </u>	<u> </u>	• •	0,5	U T	Canel		(223-1)
Education		11	12	12	06			08	373	366 ^{ab}	(1649)
ALL FIELDS		<u>07</u>	, <u>14</u>	12	10	•		06	(367)	(363 ^{ab·c})	(9375)

Note. Entries are regression and correlation coefficients without decimals. The regression coefficients tabled are for part scores contributing at least .001 to R-squared; underscored coefficients also met a .05 statistical significance criterion.

aV* significant, .05; bQ* significant, .05; cA* significant, .05

- o The regular mathematics subscore contributed at least .001 to R-squared in every analysis and is the only part score for which this was true.
- o AR and/or LR were selected as part of the most efficient part-score battery in 10 of the 12 field analyses (though with AR variance suppressed in two).

Generally speaking, the best weighted composites of selected part scores yielded somewhat higher multiple correlations with SR-UGPA than the three total test scores; no corrections for shrinkage have been made, however. In evaluating the findings in TAble 13, it is important to note that the subtests involved are of differing lengths and reliabilities, that the analysis did not attempt to adjust for these factors, and that, given moderately intercorrelated predictors such as those involved in the analysis, regression weights are sensitive to relatively small changes in validity.

Comparability of Regression Results for Unequated and Equated Total Scores

The preceding analyses were based primarily on test acores that were not equated across test forms. To what extent do patterns of findings based on unequated score data provide a basis for projecting results that might be obtained if equated part and total scores were to be employed? Table 14 presents findings bearing on the comparability of regression results for unequated (V^*, Q^*, A^*) and equated (V, Q, A) total scores on the respective tests.

While there are differences in detail in the results of the parallel analyses, the relative weighting of the verbal, quantitative, and analytical scores, and the relative magnitudes of the multiple correlation coefficients, by field, are essentially the same for the two analyses. It seems reasonable to infer that comparable results might be expected for parallel analyses involving equated and unequated part scores (see Appendix B).

From Table 14 it may be determined that the multiple correlations for the V^*, Q^*, A^* composites are somewhat lower than those for the V, Q, A composites due, it is assumed, to error associated with lack of equating for V^*, Q^* , and A^* across forms.

Summary of Trends in Findings

Major trends in the findings bearing on the predictive and/or construct validity of item-type part scores are summarized below, by test.

With respect to the verbal ability measure-

- o Althought there are some exceptions, by field, reading comprehension items (SC + RD) tend to be more valid than vocabulary items (ANT + ANA) and the same tends to be true of the RC and VO component item types.
- o RC and VO item types appear to be contributing to the prediction of



Table 14

Comparisons of Regression Results for Unequated Raw Total Scores

(V*, Q*, A*) and GRE Scaled Scores (V, Q, A), by Field

	•	Unequated Scores Beta Weights			(0)	Equated Scores Beta Weights			(B)
* •	(3)	V*	Q*	A *	(R)	. v	Q	A	(R)
. English .	(884)	. 353 ,	066	026	(402)	354	075	025	(407)
History *	(584)	345	097	-035 -	(374)	353	115	-041	(388)
Sociology	(364)	296	025	189	(442)	294	061	184	(459)
- Polit Sci	(545)	256	242	- 044	(410)	253	256	-047	(417)
Chemistry	('644)	080	281	073	(374)	. 085	294	074	(389).
Computer Sci	(647)	103	276	060	(371)	108	287	051	(377)
Mathematics	(251)	193	296	-024	(395)	217	308	-022	(425)
Elec Engin	(850)	030	317	081	(386)	040	336	075	(405)
Economics	(663)	258	147	153	(455).	272	. 159	143	(468)
Biology	(1318)	178	191	062	(350)- '	195	204	058	(370)
, Agriculture	(976)	112	177	074	(306)	125	206	066	(335)
Education	(1649)	2 2 6	149	050	(366)	224 §	149	056	(367)
All Fields	(9375)	196	187,	052	(363)	204	201	049	(377)

Note: Entries are standard partial regression (beta) weights or multiple correlation coefficients without decimals.

V*, Q*, A* are raw unequated total scores, z-scaled by form. V, Q, A are GRE scaled scores, fully equated across forms.

academic performance in fields that vary widely in apparent verbal emphasis.

o Majors in verbal fields tend to perform better on VO than on RC while majors in quantitative fields tend to perform better on RC than VO (with the anomalous exception of mathematics (see Figure 1 and related discussion).

With respect to the quantitative ability measure-

- o Data interpretation (DI) items appear to be contributing only slightly to overall predictive validity.
- o Regular mathematics (RH) items may be particularly predictive, of performance in mathematics (hypothetically, because of a greater degree of overlap between test content and curricular content for mathematica majors than for others).
- o Both RM and quantitative comparisons (QC) items appear to be contributing to prediction, though not necessarily equally so, in fields that differ widely in apparent quantitative emphasis.
- o Majors in verbal fields (for example, history, English, political science) tend to perform much better on DI items than on other quantitative item types, while the opposite is true for majors in math and science fields (for example, engineering, chemistry, computer science, mathematics).

With respect to the enalytical ability measure-

- o Based on their relative contribution to prediction of SR-UGPA, logical reasoning (LR) items appear to be underrepresented and analytical reasoning (AR) items overrepresented in the current 12-item to 38-item, LR to AR, ratio in the analytical ability measure. The shorter LR subtest appears to be as valid as the longer AR subtest.
- o Analytical reasoning items behave more like quantitative ability items while logical reasoning items behave more like verbal ability items—they may prove to be useful extensions of the two basic ability measures.
- o Majors in verbal fields perform better on LR than on AR, while the opposite is true for majors in quantitative fields; ranks of fields in terms of mean total analytical ability score differ considerably from ranks based on AR and LR means, and ranks based on AR means differ from ranks based on LR means.

Discussion

Findings regarding the GRE vocabulary and reading comprehension subtests tend to confirm and extend findings based on parallel subtests of 500



the SAT verbal seasure. These results, combined with results of factor studies indicating distinguishable "Vocabulary" and "reading comprehension" verbal, factors defined by items the those in the subtests under consideration in this study, suggest's potentially useful role for VO and RC subscores as defined for the study.

No a priori rationale was available for projecting particular patterns of validity for item-type part scores, on the quantitative and analytical However, regult congressed that the respective part ability massures. scores, are measuring somewhat different aspects of quantitative and analytical reasoning ability. Based on observed patterns of validity coefficients for quantitative subtests and average scores for different me for , the components of quantitative ability being measured by the data Nucepretation items appear to be different from those being measured by QC and RM. This is consistent with the results of a factor analysis of [sgall sets of items from the 1977 GRE Aptitude Test (Powers & Swinton, 1981) in which DI 'item gets helped to define a varimux factor called "data interpretation and technical comprehension" along with items from technical reading passages and items from the 1977 version of the analytical ability measure that seemed similar to the technical reading passages in content and style. In a factor analysis (Rock, Wertal & Grandy, 1982) that involved intercorrelations of item-type part scores paralleling those employed in this study, the loading of the DI items on the quantitative factor was less than the Wadings for QC and RM items.

The uniquely high predictive validity of regular mathematics items for mathematics majors, and evidence of differential validity for QC and RM items across fields, suggest the potential for improved assessment in separate consideration of the quantitative item types.

With respect to the analytical ability measure, perhaps the most intriguing aspect of the findings that have been reviewed is (a) the rather persistent indication that AR items tend to exhibit "quantitative" characteristics while LR items, tend to exhibit, "verbal" characteristics, and (b) that Mitems may tend to be more valid than AR items. Powers and Swinton (1981) found that logical reasoning items included in the 1977 version of the analytical ability measure were highly related to a reading. comprehension factor. And, with regard to the comparative validity of the LE and AR item types, Wild, Swinton, and Wallmark (1982, Table 22) reported that a subtest containing a 74 percent to 26 percent mix of 19 analytical and logical reasoning items was less closely related to SR-UCPA than a subtest of the same length than included only logical reasoning items (for example, r = .204 for the AR/LR combination vs SR-UGPA as compared to r = .209 for a 19-item subtest including only logical reasoning items). Combining these two frem types in a single score would appear to blunt their predictive effectiveness (see Table 9 and related discussion); moreover, the findings raise questions regarding the desirability of including more AR than LR items in the analytical ability measure since the logical reasoning items appear to have greater criterion-related validity than the analytical reasoning items.

The results that have been reviewed point up the value of evidence regarding the criterion-related validity of item types within the more

general verbal, quantitative, and analytical ability measures. Such evidence could be helpful (a) as a factor to be considered in determining the mix of items in a given ability measure—for example, in decisions regarding the proportional mix of existing item types or decisions to add or eliminate particular item types and (b) in assessments of construct validity—for example, as supplementary to the findings of factor analysis. Using data available in GRE files it would be feasible to develop, and update periodically, basic correlational results for all fields based on pooled departmental data for samples of enrolled undergraduates.*

^{*}Previous studies employing SR-UGPA as an external academic criterion (for example, Miller & Wild, 1979; Wild, Swinton, & Wallmark, 1982; Goodison & Wild, 1982) have been based on total correlation matrices (that is, test-UGPA correlations were computed for samples that were heterogeneous with respect to undergraduate department, even though homogeneous with respect to, say, broad graduate major areas). The direction and extent of covariation among means of departments on the GRE and SR-UGPA variables are not predictable—differences in mean SR-UGPA by department cannot be assumed to reflect differences in level of undergraduate performance. Accordingly, the interpretation of analyses based on total correlation matrices is complicated by the fact that such matrices include the theoretically unpredictable among—means covariances as well as the within-department covariances (see Appendic C).

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Appendix A

Supplementary Data on GRE Part Scores

Six forms of the GRE General Test were administered between October 1981 and September 1982. Table A shows (a) the number of examinees in the sample taking each form and their distribution by sex and (b) means and standard deviations of scores on selected test variables, namely, the verbal, quantitative, and analytical scaled scores, equated across forms, and the raw scores on the various item-type subtests. The latter are not equated across forms, and the average difficulty level of the items making up each subtest may vary within tests for a given form as well as across tests.

Based on the GRE scaled total scores, examinees who took forms used in the first three administrations were somewhat more able than those who took the three forms used in the last two administrations. Males constituted a majority of examinees taking certain forms, while females constituted a majority of examinees taking other forms (for example, in April and June). A majority of all members of the study sample were female.

It may be determined that the part-score means do not covary consistently with the scaled total score means, although a tendency toward positive covariation across forms between raw part scores and total scaled scores is discernible. Data not tabled indicated that the raw total scores on the respective tests covaried closely with the total scaled scores.

In z-scaling all raw scores by test form, using means and standard deviations for all examines taking each form regardless of administration date, it was assumed (a) that there would be attenuating effects on the relationship of the z-scaled scores to SR-UGPA associated with lack of equating, but (b) that those effects would be random with respect to item types across forms, and thus (c) that the relative weighting of particular item types would not be influenced by any systematic biasing effect. Evidence suggesting that these assumptions were generally valid is provided in Appendix B.



Table A. Means and Standard Deviations of Raw Part and Converted Total Scores for Selected GRE General Test Takers During 1981-82, by Test Form and Administration Dates

Form	3DGR1	anana a	On and	3		0 m a m 6		
rorm	21M+KT	3DGR2	3DCR3	• 3EGR1	K-3DGR3	3EGR2	All Forms	
Males	1584	1314	449	. 921	170	150	4588	·
Females	1497	1295	392	1116	206		4715	
Total*	3103	2633	846	2055	378	360	9375	
	, ,	2037	(/4(/	203.	3/0	300	7373	
Admins	nc&-nec	Oct-Dec-Feb	Oct-Feb	Apr	Apr	June	Total	
VARIABLE	MEAN	NE 出稿	MF A N	特差点簿	MEAN	nê.	REAN	
ANTONYNS	12.3139	13.7972	12.9741	10.4404	11.6511	11.1859		
ana: Oct	10.0140	10.5063.	11.3901	11-1086	10.6472	4.4167	12.05**	
SENT.COM	4.6210	4.44 34	10.2894	4, 7299	9.6005	9.7333	10-9336	
a fading	16.9754	* * * * * * * * * * * * * * * * * * *	14.6414	13.5173	13.0825	14,7611	*.606 *	į
VOCAD.	22.3329	27,2674	25.7191	21-4042	22.4974	20.6036	14.4521 22.5434	-40-
FO. CORP Quant. C	25.9974 24.3976	23.9537	24.9515	23.2472	23.2631	24,4944	24.5590	\$
AEG. AT.	13-0422	21.927 <u>[</u> 12.9268	21.6367	20.0442	70.3964	22.5778	21.4632	
DATA IMT	7.4170	6.7201	15.7194 6.1797	12.42 9 6 A.652è	14.5641	12.0063	17, 1972	
anal . R.	23.5203	24.0122	32.2624 '	21.4504	5.4550 21.0165	9.9020	#. \$U70	
LOG. R	7.0184	4.5047	7.4744	4.3786	4.8730	20.5250- 7.4554	23_0662	
CRE-#	528.5305	517.2541	518.6946	449. 2865	503.5485	505. 3869	0.7875	
CEE-G	588.6026	5,79.4875	605.3426	55%. 1436	558, 7564	561.5276	917.4744	
GAE-A	574.6955	569.3619	505.1655	555.2360	344.0567	554.0413	578.6725 56 7.9307	
v 4# 4#4 £	R 9 P as A a an h	SIGNACES	SICHAINS	***		***************************************		
4 704 1 2 24 2	Sicasia)	310-404101	3 : 4-4 - 1 - 1 - 2	SIGNAINS	signatus	SIGNA! 70	Signains	
anfonyms	4.0569	3.4724	7.4459	y. 4473	3.5747	3.7212	3.9913	
anal CGT	3.1747	2.4621	2.6632	3.4235	3.5477	2.7119	3.1715	
SENT,COM	2.7632	2.5547	2.5053	2.477A	2.5010	2.2148	2.6116	
PEADING VUCAR:	3,827 <i>2</i> 6,6689	9.9677 2.2688	3.7018 5.6577	3.7540	3.9661	3.9744	4.0344	
*D. COMP	e.0059	5.9404	5.377 6	6.6132 9.8499	4.0195	5.4055	0.4514	
QUANT. E	4.4364	9.0240	4,1294	4.4954	5.4AR2	6.1167	5.4651	
REG. MT.	3.0681	3.4646	3.50%	4.1444	4.9110 3.8896	5.04 04 3.4044	4. 7041	
DATA INT	1.0758	2.0520	2.2562	2.0194	2.3450	1.9204	3. 7544	
4444 . H.	6.2379	7.1577	5.7675	4.3161	. 1416	8.4946	5.0475	46
10G. #	2.1946	2.4459	2.3148	2. >/40	2.4667	2.3932	#.5942 2.4343	* •
CBE-A	114.2864	112.4384	104.4412	1:2.0135	114.0155	198, 8927	113.1023	
CRE-0	124.3348	127.4319	344,5958	129.6586	135.7218	130.6776	127.1072	
GRE-A	120.1150	126.4527	114.5545	121.2916	125.9757	120.7679	7805.181	• -

^{*}Includes individuals not identifiable by sex

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Appendix B

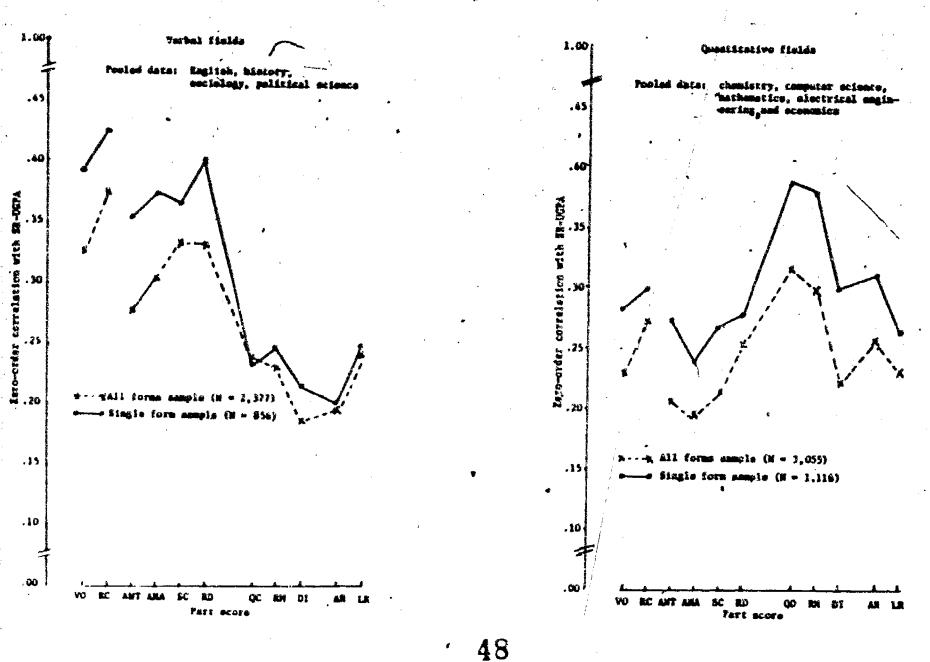
Comparability of Part-Score Validity Profiles for Single Form and Multiple Form Unequated Score Samples

Regression analyses using unequated total verbal, quantitative, and analytical scores from six different forms of the General Test and analyses employing the three GRE acaled total scores, respectively yielded entirely comparable results (see text, Table 1%, and related discussion). The relative weighting of the three total scores was consistent across analyses. As expected, the level of correlation was higher for the equated total scores than for the unequated total scores, due, it is assumed, to errors associated with lack of equating across test forms.

Parallel analyses employing equated and unequated part scores were not teasable. However, intercorrelation matrices were generated for examinees taking a single form of the test. The pattern of correlations of part scores with SR-UGPA in this sample may be compared with that for examinees taking several forms, with unequated scores, by reference to Figure 8-1.

The part-score correlational profiles for the single-form and multiple-form samples are quite similar, but the level of test-criterion correlations tends to be higher in the single-form sample than in the multiple-form sample. These results suggest strongly that conclusions regarding the Pelative criterion-related validity of various item-type part scores, based on the findings of the present study employing unequated scores, would be applicable for equated part scores.

Figure S-i. Profiles of pealed within-department correlations of CRE part scores with in-UCPA for emaininess in vertal and quantitative fields, respectively, (a) who took a single form of the CME timeral Fact and (b) who took up to six different took forms, scores not equated



Appendix C

Factors Involved in the Use of Total vs Pooled Within-Group Correlations in Validation Research

All regression analyses in this study employed pooled within-department correlation matrices. All variables were z-scaled within each department before pooling. Other research employing the self-reported UGPA as an academic criterion (for example, Miller & Wild, 1979; Wild, Swinton, & Wallmark, 1982; Goodfson & Wild, 1982) has used total correlation matrices in which coefficients were based on data for all individuals in departmentally heterogeneous samples.

Such total sample correlations are difficult to interpret because it cannot be assumed that differences in mean GPA across several departments represent substantive differences in achievement—the nature of the among means correlation between GRE scores and GPA across several departments is theoretically unpredictable since it is influenced by arbitrary differences in grading standards among departments.

Exhibit C.1 provides scatterplots of GRE verbal (or quantitative) mean and first-year graduate GPA means for samples of students from graduate departments that participated in a study of the 1977 restructured GRE General Test (Wilson, 1982).

- o Overall, the scatterplot of GRE quantitative and GPA means for chemistry, computer science, economics, and mathematics departments (upper portion of the exhibit) suggests a low positive correlation among departmental means.
- o in the lower portion of the exhibit, it may be seen that, among education departments, there is a clear tendency for mean graduate GPA to vary inversely with mean GRE score, while, for the English departments, the scatter of means suggests a generally positive, curvilinear relationship.

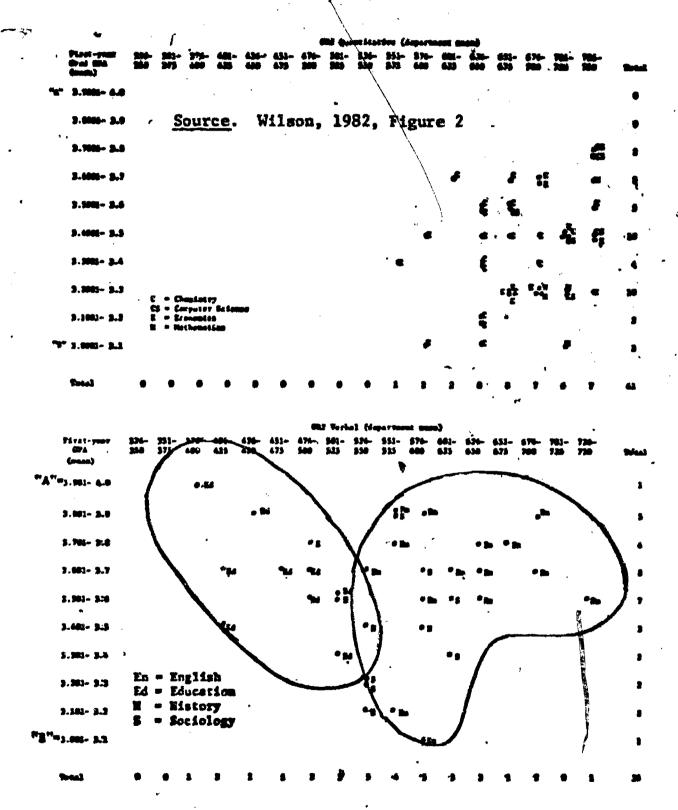
The trends illustrated in the exhibit are consistent with the proposition that neither the degree not the direction of covaration between departmental GRE and GPA means can be assumed to follow a predictable pattern. Moreover, it is reasonable to infer that the total GRE-GPA correlations for education and English majors would differ even though the pooled within-group (within-department) correlations were identical. If such were the case, the total GRE-GPA correlation should be higher for the English sample (with positive among-means correlation) than for the education sample (with negative among-means correlation).

Using data from the present study, total correlations between SR-UGPA and the respective GRE item-type part scores (prior to within-department standardization) were computed to provide a basis for comparison with the pooled within-department correlations actually used in the study. Illustrative findings are summarized in Figure C.1. Note, for example, that



Exhibit C.1

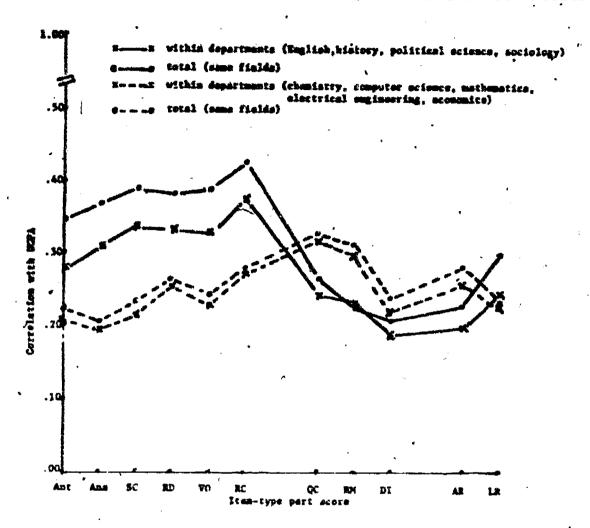
Covariation between mean GRE score and mean graduate GPA for selected departmental samples

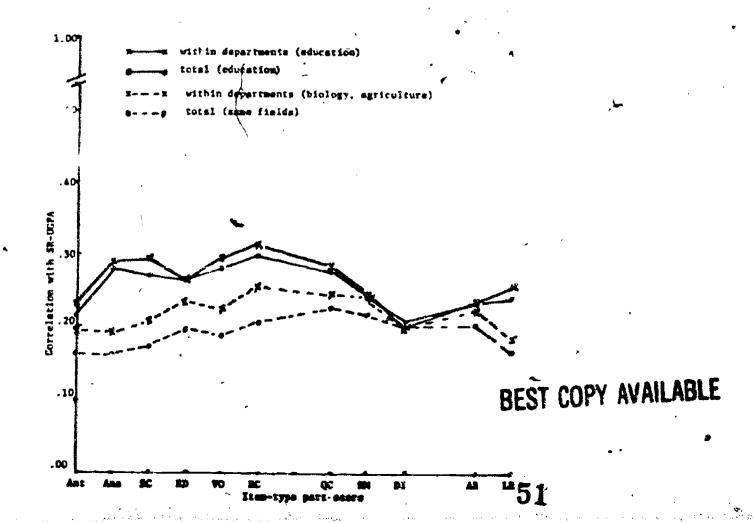


Mean GRE Aptitude Test score (GRE-V or GRE-Q as appropriate to a field) in relation to mean Tear I graduate GPA for 35 departmental samples from primarily verbal fields and 41 samples from primarily quantitative fields.



Figure C.1. Comparison of pooled within-department and total correlation between SR-UGPA and GRE scores in selected fields







in the data for education, and the combined agriculture and biology samples, the total correlation is systematically lower than the pooled within-department correlation while the opposite tends to be true for the verbal sample.

The use of total rather than within-group correlations in the present study probably would have led to somewhat_different outcomes. Conclusions regarding the relative level of validity of particular subtests for various disciplines would have been affected, for example. It is not clear whether or how outcomes bearing on the relative contribution of the various subtests to prediction of the SR-UGPA criterion might have been affected.

Strictly speaking, it would seem that the most rigorously designed studies of GRE correlations with SR-UGPA would call for the use of pooled, within-department matrices. In validation research involving GPA criteria, the use of total correlations in departmentally heterogenous samples involves elements of interpretive ambiguity that can be avoided only by using pooled within-group correlations.

